

Variable selection and feature construction using methods related to information theory

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Outline

- 1 Information Theory
 - Definitions
 - Mutual Information and Communication Channels
 - Mutual Information in Practice
- 2 Mutual Information and classification problems
 - Class Separability Measures
 - The Bayes Error
 - MI in Variable Selection
- 3 Feature Transforms based on Mutual Information
 - Maximizing Mutual Information
 - Illustrations
 - Nonlinear Transforms
 - Reducing computation
- 4 Further uses for Information Theoretic concepts
 - Learning Distance Metrics
 - Information Bottleneck
- 5 Conclusion

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Why Information Theory?

- Variables or features can be understood as a “noisy channel” that conveys information about the message
- The aim would be to select or to construct features that provide as much information as possible about the “message”
- By using information theory, variable selection and feature construction can be viewed as coding and distortion problems
- Read Shannon!

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Entropy

- Continuous random variable $X \in R^d$ representing available variables or observations and a discrete-valued random variable Y representing the class labels
- The uncertainty or entropy in drawing one sample of Y at random according to Shannon's definition:

$$H(Y) = E_y[\log_2 \frac{1}{p(y)}] = - \sum_y p(y) \log_2(p(y)). \quad (1)$$

- (Differential) entropy can also be written for a continuous variable as

$$H(X) = E_x[\log_2 \frac{1}{p(x)}] = - \int_x p(x) \log_2(p(x)) dx. \quad (2)$$

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$$H(X) = E_x[\log_2 \frac{1}{p(\mathbf{x})}] = - \int_{\mathbf{x}} p(\mathbf{x}) \log_2(p(\mathbf{x})) d\mathbf{x}. \quad (2)$$

Conditional Entropy, Mutual Information

- After having made an observation of a variable vector \mathbf{x} , the uncertainty of the class identity is defined in terms of the conditional density $p(y|\mathbf{x})$:

$$H(Y|X) = \int_{\mathbf{x}} p(\mathbf{x}) \left(- \sum_y p(y|\mathbf{x}) \log_2(p(y|\mathbf{x})) \right) d\mathbf{x}. \quad (3)$$

- Reduction in class uncertainty after having observed the variable vector \mathbf{x} is called the mutual information between X and Y

$$I(Y, X) = H(Y) - H(Y|X) \quad (4)$$

$$= \sum_y \int_{\mathbf{x}} p(y, \mathbf{x}) \log_2 \frac{p(y, \mathbf{x})}{p(y)p(\mathbf{x})} d\mathbf{x} \quad (5)$$

Same as Kullback-Leibler divergence between the joint density $p(y, \mathbf{x})$ and its factored form $p(y)p(\mathbf{x})$.

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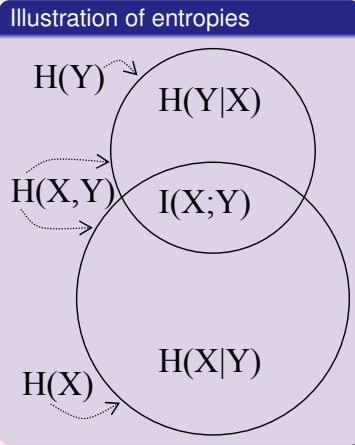
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Conditional Entropy, Mutual Information

- $H(X)$ and $H(Y)$ are each represented by a circle
- Joint entropy $H(X, Y)$ consists of the union of the circles
- Mutual information $I(X, Y)$ is the intersection of the circles
- $H(X, Y) = H(X) + H(Y) - I(X; Y)$



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Channel Coding

Shannon:

- Channel with input X and output Y'
- Rate of transmission of information $R = H(X) - H(X|Y') = I(X, Y')$
- The capacity of this particular (fixed) channel is defined as the maximum rate over all possible input distributions, $C = \max_{p(X)} R$
- Maximizing the rate = choosing an input distribution that matches the channel (under some constraints, such as fixed power or efficiency of the channel)

Analogy to variable selection and feature construction

- Real source Y is now represented (encoded) as the available variables X
- Now the channel input distribution X is fixed
- Modify how the input is communicated to the receiver by the channel either by selecting a subset of available variables or by constructing new features $\Phi = g(X, \theta)$ where g denotes some selection or construction function, and θ represents some tunable parameters
- In Shannon's case θ was fixed but X was subject to change
- "channel" capacity can be represented as $C = \max_{\theta} R$ subject to some constraints, such as keeping the dimensionality of the new feature representation as a small constant

Rate-distortion theorem

- Finding the simplest representation (in terms of bits/sec) to a continuous source signal within a given tolerable upper limit of distortion
- Would not waste the channel capacity
- Solution for a given distortion D is the representation Φ that minimizes the rate $R(D) = \min_{E(d) \leq D} I(X, \Phi)$
- Combination of the two results in a loss function

$$\mathcal{L}(p(\phi|x)) = I(X, \Phi) - \beta I(\Phi, Y). \quad (6)$$

that does not require setting constraints to the dimensionality of the representation, rather it emerges as the solution

- The representation Φ can be seen as a bottleneck that extracts relevant information about Y from X (Tishby 1999)

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Estimating Mutual Information

- Between two variables use non-parametric histogram approach (Battiti 94), but in higher dimensions any amount of data is too sparse to bin.
- Parametric class density estimates (such as Gaussians) and plug them into the definition of MI
- MI is a difference between two entropies: Entropy estimation!
 - The simplest way is the maximum likelihood estimate based on histograms
 - known to have a negative bias that can be corrected to some extent by the so-called Miller-Madow bias correction. This consists of adding $(\hat{m} - 1)/2N$ to the estimate, where \hat{m} denotes an estimate of the number of bins with nonzero probability
 - this cannot be done in many practical cases, such as when the number of bins is close to the number of observations (Paninski 93)
 - Bayesian techniques can be used if some information about the underlying probability density function is available in terms of a prior (Wolpert & Wolf 95; Zaffalon 02)

Measures other than Shannon's

- Shannon derived the entropy measure axiomatically and showed that no other measure would fulfill all the axioms
- If we want to find a distribution that minimizes/maximizes the entropy or divergence, the axioms used in deriving the measure can be relaxed and still the result of the optimization is the same distribution (Kapur, 1994)
- One example is the Renyi entropy, which is defined for a discrete variable Y and for a continuous variable X as

$$H_\alpha(Y) = \frac{1}{1-\alpha} \log_2 \sum_y p(y)^\alpha; \quad H_\alpha(X) = \frac{1}{1-\alpha} \log_2 \int_{\mathbf{x}} p(\mathbf{x})^\alpha d\mathbf{x}, \quad (7)$$

where $\alpha > 0$, $\alpha \neq 1$, and $\lim_{\alpha \rightarrow 1} H_\alpha = H$

- Quadratic Renyi entropy is straightforward to estimate from a set of samples using the Parzen window approach

Non-Parametric Estimation of Renyi Entropy

Make use of the fact, that a convolution of two Gaussians is a Gaussian, that is,

$$\int_{\mathbf{y}} G(\mathbf{y} - \mathbf{a}_i, \Sigma_1) G(\mathbf{y} - \mathbf{a}_j, \Sigma_2) d\mathbf{y} = G(\mathbf{a}_i - \mathbf{a}_j, \Sigma_1 + \Sigma_2). \quad (8)$$

Renyi entropy reduces to samplewise interactions when combined with Parzen density estimation (Principe, Fisher, and Xu, 2000).

$$\begin{aligned} H_R(Y) &= -\log \int_{\mathbf{y}} \rho(\mathbf{y})^2 d\mathbf{y} \\ &= -\log \frac{1}{N^2} \int_{\mathbf{y}} \left(\sum_{k=1}^N \sum_{j=1}^N G(\mathbf{y} - \mathbf{y}_k, \sigma^2 l) G(\mathbf{y} - \mathbf{y}_j, \sigma^2 l) \right) d\mathbf{y} \\ &= -\log \frac{1}{N^2} \sum_{k=1}^N \sum_{j=1}^N G(\mathbf{y}_k - \mathbf{y}_j, 2\sigma^2 l). \end{aligned} \quad (9)$$

Divergence Measures

- Kullback-Leibler divergence

$$K(f, g) = \int_{\mathbf{x}} f(\mathbf{x}) \log \frac{f(\mathbf{x})}{g(\mathbf{x})} d\mathbf{x} \quad (10)$$

- Variational distance (Based on the f-divergence family)

$$V(f, g) = \int_{\mathbf{x}} |f(\mathbf{x}) - g(\mathbf{x})| d\mathbf{x}. \quad (11)$$

- Quadratic divergence

$$D(f, g) = \int_{\mathbf{x}} (f(\mathbf{x}) - g(\mathbf{x}))^2 d\mathbf{x}, \quad (12)$$

- Pinsker's inequality gives a lower bound on $K(f, g) \geq \frac{1}{2} V(f, g)^2$. Since $f(\mathbf{x})$ and $g(\mathbf{x})$ are probability density functions, both are between zero and one, and $|f(\mathbf{x}) - g(\mathbf{x})| \geq (f(\mathbf{x}) - g(\mathbf{x}))^2$, and thus $V(f, g) \geq D(f, g)$. Maximizing $D(f, g)$ thus maximizes a lower bound to $K(f, g)$.

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Class Separability Measures

- 1 Sums of distances between data points of different classes.
- 2 Nonlinear functions of the distances or sums of the distances.
- 3 Probabilistic measures based on class conditional densities.
 - These measures may make an approximation to class conditional densities followed by some distance measure between densities (Battacharyya distance or divergence)
 - A Gaussian assumption usually needs to be made about the class-conditional densities to make numerical optimization tractable.
 - Equal class covariance assumption, although restrictive, leads to the well known Linear Discriminant Analysis (LDA), which has an analytic solution.
 - Some measures allow non-parametric estimation of the class conditional densities.
- 4 The Bayes error

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Relation of The Bayes Error to Mutual Information

- The Bayes risk using 0/1-loss for classification can be written as the Bayes error:

$$e_{\text{bayes}}(X) = E_x[\Pr(y \neq \hat{y})] = \int_{\mathbf{x}} p(\mathbf{x}) \left(1 - \max_i (p(y_i|\mathbf{x}))\right) d\mathbf{x}, \quad (13)$$

- An upper bound on the Bayes error (Hellman, 1970; Feder 1990)

$$e_{\text{bayes}}(X) \leq \frac{1}{2}H(Y|X) = \frac{1}{2}(H(Y) - I(Y, X)) \quad (14)$$

- A lower bound on the error also involving conditional entropy or mutual information is given by Fano's (1961) inequality

$$e_{\text{bayes}}(X) \geq 1 - \frac{I(Y, X) + \log 2}{\log(|Y|)}, \quad (15)$$

where $|Y|$ refers to the cardinality of Y .

- Both bounds are minimized when the mutual information between Y and X is maximized, or when $H(Y|X)$ is minimized.
- The bounds are relatively tight, in the sense that both inequalities can be obtained with equality

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Pairwise MI in variable selection

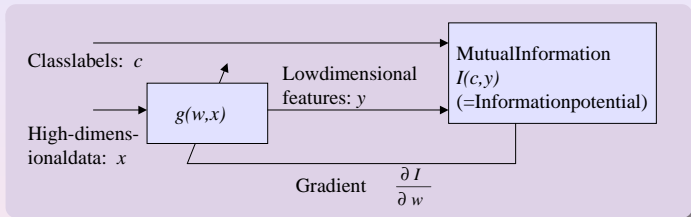
MIFS

- 1: Set $\hat{X} = \operatorname{argmax}_{X_i} I(Y, X_i)$;
 set $\Phi \leftarrow \{\hat{X}\}$;
 set $F \leftarrow \{X_1, \dots, X_N\} \setminus \{\hat{X}\}$.
- 2: For all pairs (i, j) , $X_i \in F$ and $X_j \in \Phi$
 evaluate and save $I(X_i, X_j)$ unless already saved.
- 3: Set $\hat{X} = \operatorname{argmax}_{X_i} [I(Y, X_i) - \beta \sum_{X_j \in \Phi} I(X_i, X_j)]$;
 set $\Phi \leftarrow \Phi \cup \{\hat{X}\}$;
 set $F \leftarrow F \setminus \{\hat{X}\}$,
 and repeat from step 2 until $|\Phi|$ is desired.

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Learning Feature Transforms by Maximizing Mutual Information Between Class Labels and Features



Express $I = I(\{\mathbf{y}_i, c_i\})$ in a differentiable form and perform gradient ascent (or other optimization) on \mathbf{w} , parameters of the transform g as

$$\mathbf{w}_{t+1} = \mathbf{w}_t + \eta \frac{\partial I}{\partial \mathbf{w}} = \mathbf{w}_t + \eta \sum_{i=1}^N \frac{\partial I}{\partial \mathbf{y}_i} \frac{\partial \mathbf{y}_i}{\partial \mathbf{w}}$$

1st part of the last term: **information force** that other samples exert to \mathbf{y}_i ,
 2nd part depends on the transform. If $\mathbf{y}_i = W\mathbf{x}_i$ then simply $\frac{\partial \mathbf{y}_i}{\partial W} = \mathbf{x}_i^T$.

Non-Parametric MI between Features and Labels

Labels — discrete random variable C .

Features — continuous, vector-valued Y .

Write I_T in between C and Y using the quadratic divergence:

$$\begin{aligned} I_T(C, Y) &= \sum_c \int_{\mathbf{y}} (p(c, \mathbf{y}) - p(c)p(\mathbf{y}))^2 d\mathbf{y} \\ &= \sum_c \int_{\mathbf{y}} p(c, \mathbf{y})^2 d\mathbf{y} \\ &+ \sum_c \int_{\mathbf{y}} p(c)^2 p(\mathbf{y})^2 d\mathbf{y} \\ &- 2 \sum_c \int_{\mathbf{y}} p(c, \mathbf{y}) p(c) p(\mathbf{y}) d\mathbf{y} \end{aligned} \quad (16)$$

Non-Parametric MI between Features and Labels

Using a data set of N samples and expressing class densities as their Parzen estimates with kernel width σ results in

$$\begin{aligned} I_T(\{\mathbf{y}_i, c_i\}) &= V_{IN} + V_{ALL} - 2V_{BTW} \\ &= \frac{1}{N^2} \sum_{p=1}^{N_c} \sum_{k=1}^{J_p} \sum_{l=1}^{J_p} G(\mathbf{y}_{pk} - \mathbf{y}_{pl}, 2\sigma^2 l) \\ &+ \frac{1}{N^2} \left(\sum_{p=1}^{N_c} \left(\frac{J_p}{N} \right)^2 \right) \sum_{k=1}^N \sum_{l=1}^N G(\mathbf{y}_k - \mathbf{y}_l, 2\sigma^2 l) \\ &- 2 \frac{1}{N^2} \sum_{p=1}^{N_c} \frac{J_p}{N} \sum_{j=1}^{J_p} \sum_{k=1}^N G(\mathbf{y}_{pj} - \mathbf{y}_k, 2\sigma^2 l) \end{aligned} \quad (17)$$

Gradient of the Information Potential

- First, we need the derivative of the potential, or, the force between two samples as

$$\frac{\partial}{\partial \mathbf{y}_i} G(\mathbf{y}_i - \mathbf{y}_j, 2\sigma^2 l) = G(\mathbf{y}_i - \mathbf{y}_j, 2\sigma^2 l) \frac{(\mathbf{y}_j - \mathbf{y}_i)}{2\sigma^2}. \quad (18)$$

- With this we get for V_{IN}

$$\frac{\partial}{\partial \mathbf{y}_{ci}} V_{IN} = \frac{1}{N^2 \sigma^2} \sum_{k=1}^{J_c} G(\mathbf{y}_{ck} - \mathbf{y}_{ci}, 2\sigma^2 l) (\mathbf{y}_{ck} - \mathbf{y}_{ci}). \quad (19)$$

This represents a sum of forces that other “particles” in class c exert to particle \mathbf{y}_{ci} (direction is towards \mathbf{y}_{ci}).

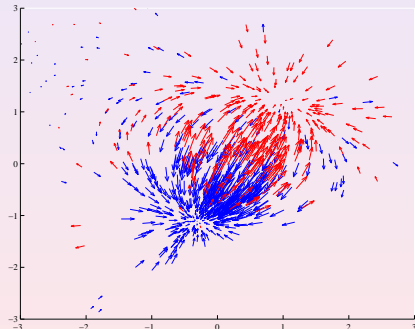
- $\frac{\partial}{\partial \mathbf{y}_i} V_{ALL}$ represents a sum of forces that other “particles” regardless of class exert to particle \mathbf{y}_{ci} (towards \mathbf{y}_i).
- The effect of $\frac{\partial}{\partial \mathbf{y}_i} V_{BTW}$ away from \mathbf{y}_{ci} , and it represents the repulsion of classes away from each other

Information Potential and Information Forces

Mutual information $I_T(\{\mathbf{y}_i, c_i\})$ can now be interpreted as an **information potential** induced by samples of data in different classes.

$\partial I / \partial \mathbf{y}_i$ can be interpreted as an **information force** that other samples exert to sample \mathbf{y}_i . It has three components:

- 1 Samples within a class attract each other
- 2 All samples attract each other
- 3 Samples between classes repel each other



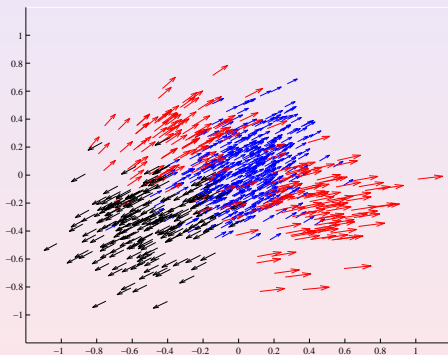
Computing $\partial I / \partial \mathbf{y}_i$ for all \mathbf{y}_i requires $O(N^2)$ operations (Torkkola, 2003).

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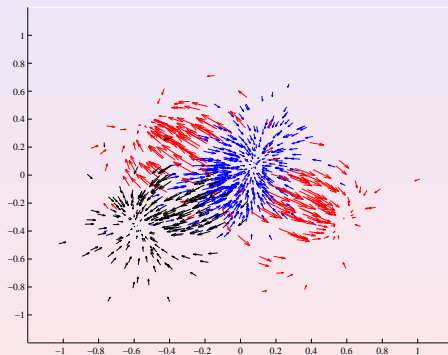
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Effect of the kernel width on the forces

Three classes in three dimensions projected onto a two-dimensional subspace.



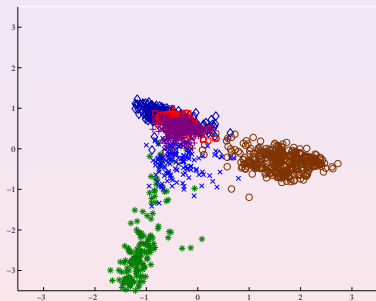
Wide kernel



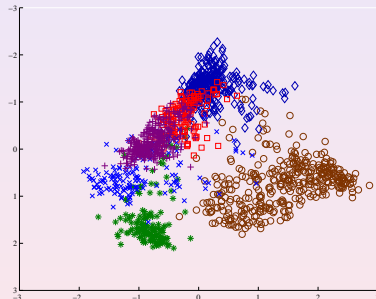
Narrow kernel

LDA vs. MMI

Landsat satellite image database from UCI repository: Six classes in 36 dimensions projected onto a two-dimensional subspace using...



LDA — compact representation of classes — on top of each other!

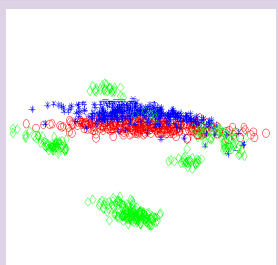


MMI — classes not as compact — need not look Gaussian — and better separated!

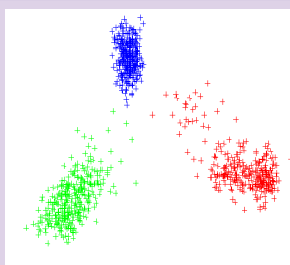
PCI vs. LDA vs. MMI

Three classes in 12 dimensions (oil pipe-flow from Aston University) projected onto a two-dimensional subspace using PCA (left), LDA or MMI with a wide kernel (middle), and MMI using a narrow kernel (right).

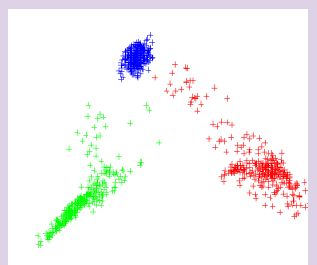
PCA



LDA/MMI wide



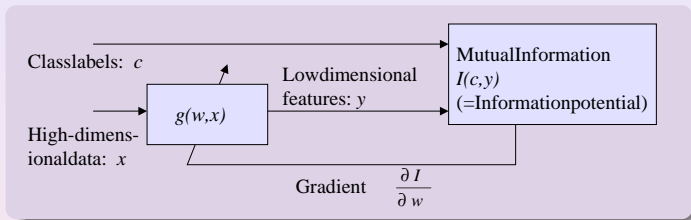
MMI narrow



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Learning nonlinear transforms using MI



- Exactly the same procedure as with linear transforms. The transform g just needs to be continuous (differentiable wrt. the parameter vector \mathbf{w}).
- The information force remains the same, the transform-dependent part $\partial \mathbf{y}_i / \partial \mathbf{w}$ will be different.

$$\mathbf{w}_{t+1} = \mathbf{w}_t + \eta \frac{\partial I}{\partial \mathbf{w}} = \mathbf{w}_t + \eta \sum_{i=1}^N \frac{\partial I}{\partial \mathbf{y}_i} \frac{\partial \mathbf{y}_i}{\partial \mathbf{w}}$$

Nonlinear transforms

Multilayer perceptrons

- Gradient of the output w.r.t. weights using (information) backpropagation
- Hidden layer activation - tanh
- Output layer activation:
 - Linear: Need orthonormalized weights in output layer
 - Tanh: Can use data-independent kernel width
- Weight initialization, partially by LDA

Radial basis function networks

- Basis functions by EM separately for each class as mixtures of diagonal-covariance Gaussians
- Two options:
 - Use MMI only to learn the output layer (this work)
 - Learn all the parameters using MMI

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Stochastic gradient

- Instead of interactions between all pairs of data points, take a sample of just two.
- Samples of the same class: NO update! Samples in different classes:

$$\begin{aligned}
 W_{t+1} &= W_t + \eta \sum_{i=1,2} \frac{\partial l}{\partial \mathbf{y}_i} \frac{\partial \mathbf{y}_i}{\partial W} \\
 &= W_t - \frac{\eta}{8\sigma^2} G(\mathbf{y}_1 - \mathbf{y}_2, 2\sigma^2 l) (\mathbf{y}_2 - \mathbf{y}_1) (\mathbf{x}_1^T - \mathbf{x}_2^T) \quad (20)
 \end{aligned}$$

- Full gradient using all pairs, and stochastic gradient using just one pair at a time are two ends of a spectrum: It is more desirable to take as large a random sample of the whole data set as possible, and to compute all the mutual interactions between those samples for one update of W .

Semi-Parametric Density Estimation

- Construct a Gaussian Mixture Models model in the low-dimensional output space after a random or an informed guess as the transform:

$$p(\mathbf{y}|c_p) = \sum_{j=1}^{K_p} h_{pj} G(\mathbf{y} - \mathbf{m}_{pj}, S_{pj}) \quad (21)$$

- The same samples are used to construct a GMM in the input space using the *same exact assignments of samples to mixture components* as the output space GMMs have. Running the EM-algorithm in the input space is now unnecessary since we know which samples belong to which mixture components.
- Now we have GMMs in both spaces and a transform mapping between the two
- Avoid operating in the high-dimensional input space altogether

Video clips

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Fisher Information

Options to make relevant information more explicit:

- 1 Variable selection.
- 2 Feature construction.

Selection/construction matrix defines a **global** Euclidean metric

$$d_A^2(\mathbf{x}, \mathbf{x}') = (\mathbf{x} - \mathbf{x}')^T S^T S (\mathbf{x} - \mathbf{x}') = (\mathbf{x} - \mathbf{x}')^T A (\mathbf{x} - \mathbf{x}')$$

- 3 Learning a distance metric locally relevant to target (or some auxiliary variable): $A = A(\mathbf{x})$

$$d_A^2(\mathbf{x}, \mathbf{x} + d\mathbf{x}) = d\mathbf{x}^T A(\mathbf{x}) d\mathbf{x}.$$

Metric should reflect the divergence between conditional distributions of the target.

$$d_J^2(\mathbf{x}, \mathbf{x} + d\mathbf{x}) = D_{KL}[p(y|\mathbf{x})||p(y|\mathbf{x} + d\mathbf{x})] = \frac{1}{2} d\mathbf{x}^T J(\mathbf{x}) d\mathbf{x}.$$

Embedded into e.g. a clustering algorithm results in "semi-supervised" clustering that reflects the auxiliary variable (Peltonen, 2004)

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Information Bottleneck

If X are the original data, Φ is a sought representation, and Y variable(s) of importance, minimizing loss function

$$\mathcal{L}(p(\phi|x)) = I(X, \Phi) - \beta I(\Phi, Y)$$

leads to

IB solution

- 1 $p(\phi|x) = \frac{p(\phi)}{Z(\beta, x)} \exp(-\beta D_{KL}[p(y|x), p(y|\phi)])$
- 2 $p(y|\phi) = \frac{1}{p(\phi)} \sum_x p(y|x)p(\phi|x)p(x)$
- 3 $p(\phi) = \sum_x p(\phi|x)p(x)$

For example, if X are documents, Y are words, and Φ are word clusters, probability of cluster membership decays exponentially with KL-divergence between the word distributions in document x and cluster ϕ (Tishby, 1999)

Conclusion

- Shannon's seminal work showed how mutual information provides a measure of the maximum transmission rate of information through a channel
- Analogy to variable selection and feature construction with mutual information as the criterion to provide maximal information about the variable of interest
- Best thing since sliced bread for variable selection / feature construction?
- Maybe not - estimation problems from high-dimensional small(ish) data sets

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